

An Analysis of the Determinants and Factors of Physical Education Attendance in the Fourth Quarter

Introduction

This project attempted to gain a more complete understanding of the internal and external factors that drive or deter 4th quarter attendance in Physical Education classes.

Data

The data was obtained from the sports center, in the form of attendance sheets and class rosters from the spring of 2004-2009 that contained information on the students themselves such as sex, status towards graduation and their major as well as gleaned class characteristics such as class size, time commitment, time of meeting, gender percentage, intensity of the activity in the class. Weather data was also obtained from the National Climatic Data Center, specifically the measuring station used was the Portland International Airport.

Variable	Obs	Mean	Std. Dev.	Min	Max
timer	837	4.373955	1.427663	1	6
attendrate	837	.5997544	.3508946	0	1
intensity	837	3.188769	.7467429	1	4
timecom	837	109.1398	24.6502	80	220
bf4	837	.3225806	.4677433	0	1
gender	837	.3357228	.4725249	0	1
status	837	2.139785	1.067487	1	4
classsize	837	14.67025	8.596134	1	41
cornc	837	.5746714	.4946883	0	1
genratio	837	.3369295	.2968043	0	1
isUnd	837	.1911589	.393449	0	1
prcp	837	23.20704	6.879986	12.65	41.55
avghigh	837	149.8975	13.94548	131.6098	187.65
Major	0				
ClassID	0				
ltc	837	4.672691	.1890199	4.382027	5.393628

Challenges

Though all the classes in the previously mentioned time frame contain information about whether people received credit, around 40% of the classes don't have

information concerning the students themselves or the days they attended. I concluded that this would not be debilitating hindrance in attempting to answer the question because there is enough data to draw conclusion on a random eclectic mix of classes.

Analysis

OLS Explorations

The large t-statistic of the credit/no credit variable is certainly a red flag concerning its admittedly special relationship with the dependent variable.

```
reg attendrate timecom intensity gender status prcp avghigh cornc
```

Source	SS	df	MS	Number of obs = 837		
Model	73.8559378	7	10.5508483	F(7, 829) = 300.80		
Residual	29.0782567	829	.035076305	Prob > F = 0.0000		
				R-squared = 0.7175		
				Adj R-squared = 0.7151		
Total	102.934195	836	.123127027	Root MSE = .18729		

attendrate	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
timecom	-.0008133	.000291	-2.79	0.005	-.0013844	-.0002421
intensity	-.0146804	.0097521	-1.51	0.133	-.0338221	.0044613
gender	.0196603	.0140024	1.40	0.161	-.007824	.0471445
status	-.012279	.0061144	-2.01	0.045	-.0242806	-.0002774
prcp	-.0001472	.0009601	-0.15	0.878	-.0020317	.0017373
avghigh	.0013093	.0004772	2.74	0.006	.0003725	.002246
cornc	.5817143	.0134765	43.17	0.000	.5552623	.6081664
_cons	.2278634	.0873895	2.61	0.009	.0563327	.3993942

If one chooses to omit it one finds that similar variables stay statistically significant but the R^2 value becomes very small indicating poor explanatory power

```
reg attendrate intensity timecom bf4 gender status classsize genratio isUnd prcp avghigh
```

Source	SS	df	MS	Number of obs = 837		
Model	8.98703207	10	.898703207	F(10, 826) = 7.90		
Residual	93.9471625	826	.113737485	Prob > F = 0.0000		
				R-squared = 0.0873		
				Adj R-squared = 0.0763		
Total	102.934195	836	.123127027	Root MSE = .33725		

attendrate	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
intensity	-.0172774	.0193461	-0.89	0.372	-.0552507	.0206958
timecom	-.001294	.0005447	-2.38	0.018	-.002363	-.0002249
bf4	-.0098879	.0276266	-0.36	0.720	-.0641146	.0443387
gender	.0434626	.0316177	1.37	0.170	-.0185978	.105523
status	-.0238113	.0118972	-2.00	0.046	-.0471636	-.0004591
classsize	-.0025306	.0017502	-1.45	0.149	-.005966	.0009049
genratio	.0487809	.0551416	0.88	0.377	-.0594533	.1570151
isUnd	.0019464	.0316236	0.06	0.951	-.0601256	.0640184
prcp	.0033157	.0017742	1.87	0.062	-.0001668	.0067982

avghigh		.0046773	.0008791	5.32	0.000	.0029518	.0064027
_cons		.0778785	.1738487	0.45	0.654	-.2633587	.4191157

However, if one also notices the significance of the constant term, it also appears dubious. Dropping the constant term yields a much more attractive R^2 value.

```
reg attendrate intensity timecom bf4 gender status classsize genratio isUnd pr
> cp avghigh, noconstant
```

Source	SS	df	MS	Number of obs =	837
Model	310.03762	10	31.003762	F(10, 827) =	272.85
Residual	93.9699867	827	.113627553	Prob > F =	0.0000
				R-squared =	0.7674
				Adj R-squared =	0.7646
Total	404.007606	837	.482685312	Root MSE =	.33709

attendrate	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
intensity	-.0162248	.0191935	-0.85	0.398	-.0538985 .021449
timecom	-.0012203	.000519	-2.35	0.019	-.0022389 -.0002016
bf4	-.0107228	.0275504	-0.39	0.697	-.0647997 .0433541
gender	.0434083	.0316022	1.37	0.170	-.0186215 .1054382
status	-.0229114	.0117207	-1.95	0.051	-.0459172 .0000944
classsize	-.0022384	.0016234	-1.38	0.168	-.0054249 .000948
genratio	.05517	.0532394	1.04	0.300	-.0493303 .1596703
isUnd	.0037185	.03136	0.12	0.906	-.0578361 .0652731
prcp	.0036098	.0016475	2.19	0.029	.0003761 .0068436
avghigh	.0050167	.0004453	11.27	0.000	.0041428 .0058907

If one narrows the model specification search further we can see that we are at a bit of an impasse.

```
reg attendrate timecom status prcp avghigh, noconstant
```

Source	SS	df	MS	Number of obs =	837
Model	307.977667	4	76.9944167	F(4, 833) =	667.88
Residual	96.0299396	833	.11528204	Prob > F =	0.0000
				R-squared =	0.7623
				Adj R-squared =	0.7612
Total	404.007606	837	.482685312	Root MSE =	.33953

attendrate	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
timecom	-.0016631	.0004391	-3.79	0.000	-.002525 -.0008012
status	-.0262258	.0108604	-2.41	0.016	-.0475428 -.0049088
prcp	.0041404	.0016375	2.53	0.012	.0009263 .0073544
avghigh	.0049462	.0003789	13.05	0.000	.0042024 .0056899

```
.
.
. estimates store sub

. reg attendrate timecom status avghigh cornc
```

Source	SS	df	MS	Number of obs =	837
Model	73.6766459	4	18.4191615	F(4, 832) =	523.79
Residual	29.2575486	832	.035165323	Prob > F =	0.0000
				R-squared =	0.7158
				Adj R-squared =	0.7144
Total	102.934195	836	.123127027	Root MSE =	.18752

attendrate	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
timecom	-.0010106	.0002651	-3.81	0.000	-.0015309	-.0004903
status	-.0128664	.0060834	-2.11	0.035	-.024807	-.0009257
avghigh	.0013095	.000475	2.76	0.006	.0003771	.0022419
cornc	.5847748	.0133512	43.80	0.000	.5585689	.6109807
_cons	.2052378	.0802179	2.56	0.011	.0477846	.362691

. estimates store full

. estimates stats full sub

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
full	837	-310.5868	215.8662	5	-421.7325	-398.0833
sub	837	.	-281.5304	4	571.0608	589.9801

Using only the statistically significant variables in each case leads to similar R^2 values if we look at the information criteria it appears that we'd have to make a choice between poor fit evidenced by negative Akiake & Bayesian values with the credit variable or lots of complexity evidenced by the large values of the two in the model with the omitted credit variable and dropped constant. Thankfully we are not forced to just rely on OLS.

A Problematic Variable

We are bound to encounter interesting phenomena with the credit/no credit variable. In practice instructors decide whether to award credit based on their degree of attendance which means that the dependent variable exogenously determines the credit or no credit variable. However, students can possibly know within three classes that they will not receive credit and still choose to come for reasons such as enjoyment or necessity or what have you. This may make the credit variable correlated with the error term.

Application of Two Stage Least Squares

After performing a Hausman test of endogeneity of cornc, it has been shown that there is collinearity between vhat, the error of the first regression with cornc being the dependent variable and the variable attendance rate. Dropping the cornc term from the second equation shoes that vhat is markedly statistically significant from zero.

It would probably not all that wise to use the two stage least squares at this time because any instrument we'd potentially use would need to not be correlated with the y of our

scenario and given the overwhelming collinearity that cornc has on the attendance rate, such an outcome is unlikely

Probit & Logit

In applying the probit and logit model to the data, it became increasingly clear that it would be more appropriate to use a probit model to better capture the relationship between attendance and the other explanatory variables. If we are to use attendance rate as an explanatory variable for credit/no credit we get a very high pseudo r-squared values with the probit model, however if we use the logit model and use credit/no credit as a predict stata drops the variable because it successfully predicts a high attendance rate when they receive credit which is essentially a statistical tautology in this case and doesn't tell us much else about it's predictive ability with the rest of the data.

```
logit attendrate cornc timecom gender classsize prcp avghigh
```

```
note: cornc != 0 predicts success perfectly
      cornc dropped and 481 obs not used
```

```
Iteration 0:   log likelihood = -228.18209
Iteration 1:   log likelihood = -205.35721
Iteration 2:   log likelihood = -204.45754
Iteration 3:   log likelihood = -204.45204
Iteration 4:   log likelihood = -204.45204
```

Logistic regression	Number of obs	=	356
	LR chi2(5)	=	47.46
	Prob > chi2	=	0.0000
Log likelihood = -204.45204	Pseudo R2	=	0.1040

attendrate	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
cornc	0	(omitted)				
timecom	-.0060344	.0043628	-1.38	0.167	-.0145854	.0025165
gender	-.4549471	.2753227	-1.65	0.098	-.9945696	.0846754
classsize	-.0300788	.0142009	-2.12	0.034	-.0579121	-.0022455
prcp	.1143427	.028376	4.03	0.000	.0587268	.1699585
avghigh	.0217899	.008936	2.44	0.015	.0042756	.0393042
_cons	-3.720066	1.530148	-2.43	0.015	-6.719102	-.7210307

Further the probit model was able to illustrate a model of good fit even without an extremely collinear variable probably because such a variable was not very applicable.

```
probit attendrate status timecom classsize prcp avghigh
```

```
Iteration 0:   log likelihood = -345.81709
Iteration 1:   log likelihood = -308.80956
Iteration 2:   log likelihood = -307.65697
Iteration 3:   log likelihood = -307.65449
Iteration 4:   log likelihood = -307.65449
```

Probit regression	Number of obs	=	837
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```

Log likelihood = -307.65449
LR chi2(5)      =      76.33
Prob > chi2     =      0.0000
Pseudo R2      =      0.1104

```

attendrate	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
status	-.0827628	.0537173	-1.54	0.123	-.1880468	.0225211
timecom	-.0044244	.0020962	-2.11	0.035	-.0085328	-.000316
classsize	-.0127608	.0061527	-2.07	0.038	-.0248198	-.0007018
prcp	.0539664	.0121431	4.44	0.000	.0301663	.0777664
avghigh	.0193205	.0040677	4.75	0.000	.0113479	.027293
_cons	-2.113469	.7194963	-2.94	0.003	-3.523656	-.7032828

In this particular model it made sense not to drop status because the loss in an insignificant variable would lead to comparatively steep drop in R-squared value.

Discussion & Conclusion

If one examines the OLS observations and the Probit model observations we can conclude that status towards graduation, time commitment, precipitation and average high temperature throughout spring have similar influences on the attendance rates of Reedies in fourth quarter pe classes overall. If one chooses to look at results by major, the results are a bit more specific:

```
probit attendrate timecom classsize prcp avghigh if Major == "ENG"
```

```

Probit regression
Log likelihood = -31.995264
Number of obs   =      102
LR chi2(4)      =      9.90
Prob > chi2     =      0.0421
Pseudo R2      =      0.1340

```

attendrate	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
timecom	-.0110808	.0055558	-1.99	0.046	-.0219699	-.0001918
classsize	-.025218	.0233303	-1.08	0.280	-.0709445	.0205084
prcp	.0164544	.038404	0.43	0.668	-.0588161	.0917249
avghigh	.0218337	.0166862	1.31	0.191	-.0108706	.054538
_cons	-.6960805	2.640665	-0.26	0.792	-5.871689	4.479528

```
probit attendrate timecom classsize prcp avghigh if Major == "BIOL"
```

```

Probit regression
Log likelihood = -26.011143
Number of obs   =      61
LR chi2(4)      =     11.18
Prob > chi2     =      0.0246
Pseudo R2      =      0.1769

```

attendrate	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
timecom	-.0333493	.012573	-2.65	0.008	-.057992	-.0087067
classsize	.0114838	.0201506	0.57	0.569	-.0280107	.0509782
prcp	.0843066	.0387896	2.17	0.030	.0082803	.160333
avghigh	.0010203	.0151189	0.07	0.946	-.0286122	.0306528
_cons	2.085181	2.885376	0.72	0.470	-3.570051	7.740413

```
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probit attendrate timecom classsize prcp avghigh if Major == "ANTH"
```

```
Probit regression                                Number of obs   =          48
                                                LR chi2(4)      =         14.36
                                                Prob > chi2     =         0.0062
Log likelihood = -18.654384                    Pseudo R2      =         0.2780
```

attendrate	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
timecom	-.0121093	.0073521	-1.65	0.100	-.0265192	.0023006
classsize	-.0649974	.0299915	-2.17	0.030	-.1237797	-.0062151
prcp	.0399769	.0523738	0.76	0.445	-.0626739	.1426276
avghigh	.0169272	.0216199	0.78	0.434	-.025447	.0593014
_cons	.0577167	3.399837	0.02	0.986	-6.605842	6.721275

While some of the effects carry over to other majors, it was difficult to achieve a comprehensive understanding of major effects in part because even in pe enrollment there is a noticeable pre-ponderance of some majors over others and we can expect to see the coefficient estimates for English more closely match the total population's estimates because they form the largest contingent of students in 5 years of classes. It is rather hard understand and demonstrate the major effects of general literature.

Additionally females make up 556 of the 837 observations. Which may mean that the estimates of the effects may only be reflective of classes that are popular with females. But the fact that gender hasn't been implicated in any statistically significant sense in any model is encouraging given that I have data for classes where guys predominate or are in equal portion.

Omitted Variable Bias

This was probably the biggest problem lurking in the background of this project, I was unable to get specific dates of Qualls and Hum Papers and thesis due dates so I scaled down the specificity of the potential time data I'd look at. As a result, the hand-wavy solution was to hope that the status variable would take up the slack in determining how differing external factors affected upper and underclassmen which is practically an invitation to inefficient and biased estimators. This particular approach can say nothing of which class is worse in terms of hardship that would preclude going to attend p.e. Further, the weather variables were also averaged out in order to apply them in a blanket like fashion over each quarter instead of each day or week.

Heteroskedasticity

Also another problem especially in the probit model in that it causes the maximum likelihood estimator to inefficient. Greater error that occurs at different levels

is a large problem in that the coefficients are constrained in order to fit a function that behaves close to asymptotically.

Autocorrelation

I did not see this as being much of a problem considering that the variables themselves seemed to have such little explanatory power right off the bat although they were found to be significant.

In conclusion, I'd posit that it would likely be better to examine students on a completely daily or weekly basis with weather being the changing explanatory variable and the class attributes being the unchanging dummy variables. Given that this approach would make more of the data actually usable, there would be more degrees of freedom and more possible classes with different attributes to use.