

Econ 312

Monday, March 30 Estimation of time-series regression models with stationary variables

Reading: Online time-series Chapter 2 Class notes: Pages 101 to 105 Daily problem: Breusch-Godfrey test (assigned for 3/16)

Today's Far Side offering





Context and overview

- Last class: We talked in detail about stationarity and its implications for time-series regression
- **Today**: We discuss how to proceed with estimation if all variables are stationary
 - Assumptions underlying use of OLS
 - Autocorrelation in the error
 - Implications
 - Testing
 - Correcting

Exogeneity and strict exogeneity

- To use OLS, we need the independent variable to be **stationary** and **exogenous**
- Exogeneity is the standard OLS assumption that the error is uncorrelated with *x*: $E(u_t | x_t) = 0$
- In a time-series context, we also need it to be uncorrelated with past and perhaps future *x*
- **Exogeneity**: $E(u_t | x_t, x_{t-1}, ...) = 0$
- Strict exogeneity: $E(u_t | ..., x_{t+2}, x_{t+1}, x_t, x_{t-1}, x_{t-2}, ...) = 0$

TS assumptions for OLS

- TS1: Linear model with additive error
- TS2: No perfect multicollinearity
- **TS3**: Regressor is **exogenous** to u^{*}
- TS4: Homoskedasticity
- **TS5**: No autocorrelation between one u_t and another u_{t-s}
- **TS6**: Error term is **normally distributed** (not totally necessary)

Autocorrelation in x and u

- What's the distinction between *x* and *u*?
 - We observe *x* but not *u*, but both affect *y*
- Nearly all time-series variables are autocorrelated over time
 - That includes the unobserved variables that we put in *u*
 - It also includes the right-hand variables *x*
- Autocorrelation in *x* does not violate the TS assumptions that underlie OLS estimation
- Autocorrelation in *u* violates assumption TS-5, so it requires attention in choosing and evaluating estimators

Serial autocorrelation

- Assumption TS-5 is almost never satisfied
- **Implications** ~ heteroskedasticity:
 - OLS is unbiased and consistent
 - OLS is not fully efficient
 - OLS standard errors are biased, so we can't use usual *t* statistics

• Solutions

- Estimate by OLS but use HAC robust standard errors
- Estimate by efficient generalized least squares
- In practice, can often reduce autocorrelation by **adding more lags** to achieve a **dynamically complete model**



Detecting serial correlation

- OLS estimator is consistent, so can use **OLS residuals** as basis for testing for autocorrelation
- We would be looking to see if \hat{u}_t is correlated with \hat{u}_{t-s} (for s = 1 in particular)
 - Recall that we defined $\operatorname{corr}(u_t, u_{t-s}) = \rho_s$
 - We want to test H_0 : $\rho_s = 0$
 - An obvious choice would be to look at the sample autocorrelations r_s
 - Three common tests are tests of correlation in OLS residuals



Breusch-Godfrey test

- Regress either y_t or \hat{u}_t on all of your x variables, plus one or more lags of the residuals: $\hat{u}_{t-1}, \hat{u}_{t-2}, \dots, \hat{u}_{t-s}$
- If there is no autocorrelation, the coefficients on the lagged residuals are zero
- Test the significance of the lagged residuals with
 - Joint *F* test, if using *y* as the dependent variable (*t* statistic works if s = 1)
 - Lagrange multiplier test: $T \times R^2 \sim \chi^2$ with *s* degrees of freedom, if using residuals, where *T* is the number of observations

Other tests

- Page 102 of the notes describes two other tests that are less common:
- Durbin-Watson test was developed decades ago, but is awkward to use
 - You'll see it in older papers
 - The critical values for the test depend on the sample, so only upper and lower bounds are tabulated
- **Box-Lyung** *Q* **test** (sometimes called the "portmanteau test") is sometimes used
 - Formula is on page 102 of the class notes



Newey-West (HAC consistent) standard errors

- If inefficient OLS estimators are acceptable, we can conduct valid tests using **standard errors that are robust** to autocorrelation (and heteroskedasticity) developed by **Newey and West**
- The formula is complicated (see pages 103 and 104) and there is one parameter that must be specified: the **number of lags** to use in approximating the standard errors
 - This is independent of the number of lags in the actual model
 - Stock and Watson suggest using $m = 0.75T^{1/3}$
 - Implement in Stata with the post-estimation command newey, lags(m)



GLS estimation with AR(1) error

• Suppose that

 $y_t = \beta_0 + \beta_1 x_t + u_t$ $u_t = \rho u_{t-1} + \varepsilon_t, \text{ with } \varepsilon \text{ being white noise and } -1 < \rho < 1$

- As when we used weighted least squares to solve heteroskedasticity, we seek to transform the model into one with a classical error term
- In this case, we can "**quasi-difference**" *y* and *x* by subtracting their lagged values multiplied by ρ (assumed to be known for now)



Quasi-differencing

$$\tilde{y}_{t} = \begin{cases} y_{t}\sqrt{(1-\rho^{2})}, & t=1, \\ y_{t}-\rho y_{t-1}, & t=2,3,...,T, \end{cases} \quad \tilde{x}_{t} = \begin{cases} x_{t}\sqrt{(1-\rho^{2})}, & t=1, \\ x_{t}-\rho x_{t-1}, & t=2,3,...,T, \end{cases} \quad \tilde{u}_{t} = \begin{cases} u_{t}\sqrt{(1-\rho^{2})}, & t=1, \\ u_{t}-\rho u_{t-1}, & t=2,3,...,T. \end{cases}$$

- We can then estimate the model $\tilde{y}_t = (1-\rho)\beta_0 + \beta_1\tilde{x}_t + \tilde{u}_t$
 - This model has an error term that is not serially correlated
 - If TS-5 was the only OLS assumption violated, then OLS is fully efficient for the quasi-differenced model because all assumptions are satisfied
- Note the asymmetric treatment of the first observation: some methods simply delete this observation
- What is ρ ? How about using $\hat{\rho} = \operatorname{corr}(\hat{u}_t, \hat{u}_{t-1})$?

Estimating $\boldsymbol{\rho}$ with OLS residuals

• **Two-step** procedure

- 1. Run OLS regression and get residuals; estimate $\hat{\rho} = \operatorname{corr}(\hat{u}_t, \hat{u}_{t-1})$
- 2. Quasi-difference the model and run OLS to get efficient estimator
- **Prais-Winsten** method uses all *T* observations
- Cochrane-Orcutt method omits the first observation
- In Stata, use the **prais** command (option **corc** to get Cochrane-Orcutt)



A problematic special case

- OLS residuals are consistent provided the error is uncorrelated with *x*, but there is an important case where this is violated
- Lagged dependent variable: If y_{t-1} is a regressor, that regressor is affected by lagged error u_{t-1}
 - If *u* is autocorrelated, then u_{t-1} also affects u_t , making the regressor y_{t-1} correlated with u_t
 - TS-3 is violated, making OLS biased and inconsistent
 - Can't rely on OLS residuals to provide estimate of ρ
 - Hildreth-Lu procedure searches directly for the $\hat{\rho}$ that minimizes the model SSR rather than using (inconsistent) OLS residuals
 - Use ssesearch option in prais to do use Hildreth-Lu



Review and summary

- Exogeneity and strict exogeneity are important in time series regressions
- The assumptions underlying OLS are slightly modified
- The error term is nearly always serially correlated in time-series regressions
- This leads OLS to be inefficient and its standard errors to be biased
- We can use Newey-West HAC robust standard errors
- We can use GLS to get efficient estimators that are better than OLS

From The Devil's Dictionary

Distance, *n*. The only thing that the rich are willing for the poor to call theirs, and keep.

[Note that for Ambrose Bierce, as well as for me, "distance" is a NOUN, not a VERB. I refuse to engage in "social distancing," though I am perfectly willing to be "keeping social distance"!]



What's next?

- This session has examined the problem of a serially correlated error term in a time-series regression model
- In the next class (April 1), we consider how to estimate models in which *y* responds to *x* slowly over time, so that lagged values of *x* affect the current value of *y*
 - These are called **distributed-lag models**
 - The lags here relate to the **deterministic** part of the model, not to the error term
 - These models are covered in Chapter 3 of the online readings